

# 5-2 Two Continuous Random Variables

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## 5-2.1 Joint Probability Distribution

### Definition

A **joint probability density function** for the continuous random variables  $X$  and  $Y$ , denoted as  $f_{XY}(x, y)$ , satisfies the following properties:

(1)  $f_{XY}(x, y) \geq 0$  for all  $x, y$

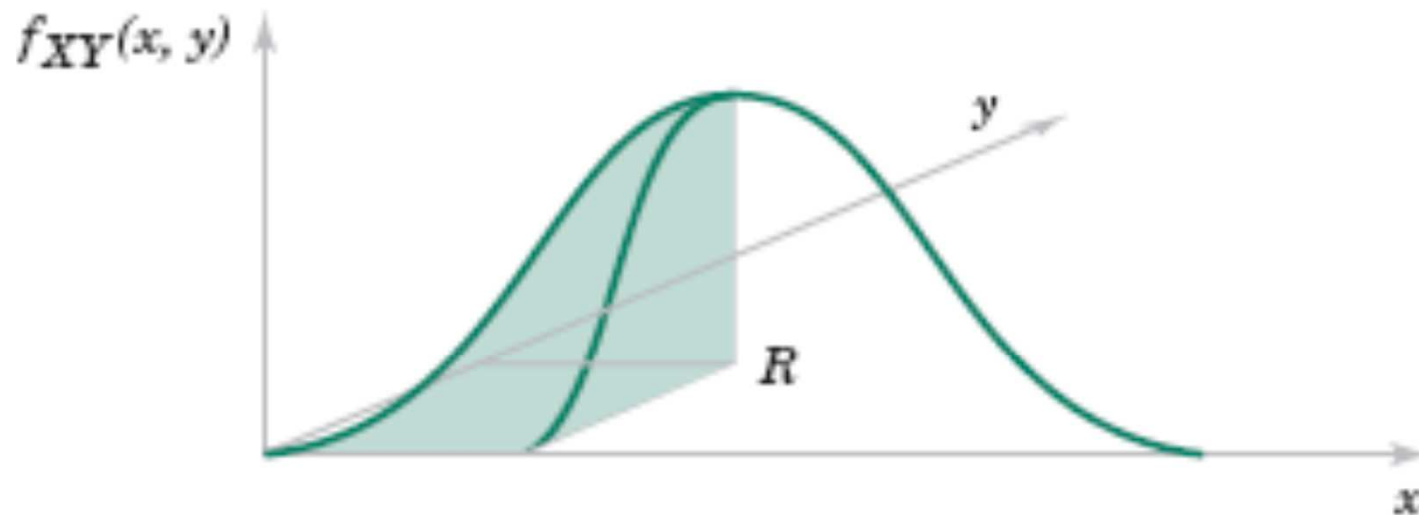
(2) 
$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{XY}(x, y) dx dy = 1$$

(3) For any region  $R$  of two-dimensional space

$$P((X, Y) \in R) = \iint_R f_{XY}(x, y) dx dy \quad (5-14)$$

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Probability that  $(X, Y)$  is in the region  $R$  is determined by the volume of  $f_{XY}(x, y)$  over the region  $R$ .

**Figure 5-6** Joint probability density function for random variables  $X$  and  $Y$ .

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### Example 5-12

Let the random variable  $X$  denote the time until a computer server connects to your machine (in milliseconds), and let  $Y$  denote the time until the server authorizes you as a valid user (in milliseconds). Each of these random variables measures the wait from a common starting time and  $X < Y$ . Assume that the joint probability density function for  $X$  and  $Y$  is

$$f_{XY}(x, y) = 6 \times 10^{-6} \exp(-0.001x - 0.002y) \quad \text{for } x < y$$

Reasonable assumptions can be used to develop such a distribution, but for now, our focus is only on the joint probability density function.

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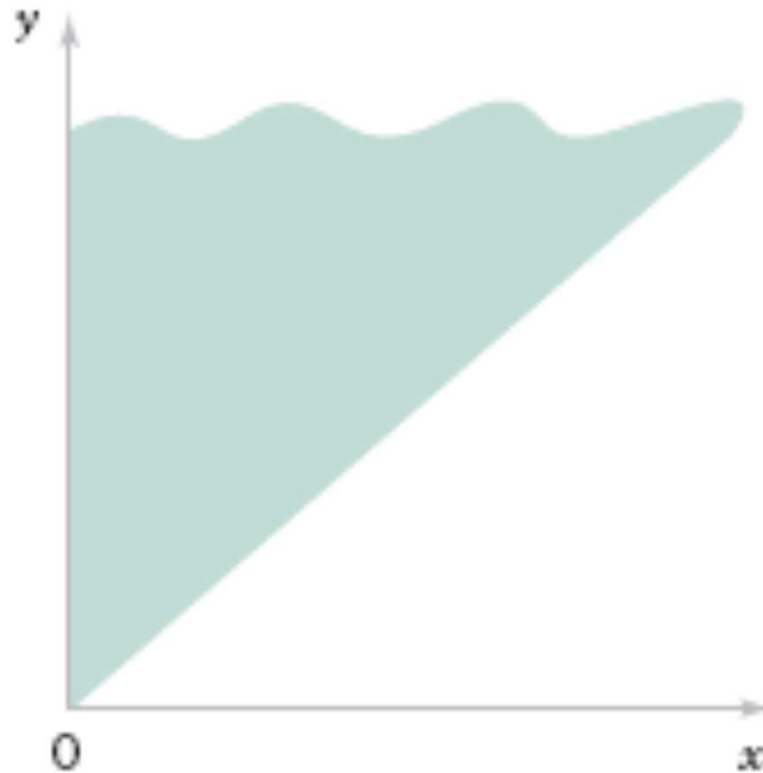
### Example 5-12

The region with nonzero probability is shaded in Fig. 5-8. The property that this joint probability density function integrates to 1 can be verified by the integral of  $f_{XY}(x, y)$  over this region as follows:

$$\begin{aligned}\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{XY}(x, y) dy dx &= \int_0^{\infty} \left( \int_x^{\infty} 6 \times 10^{-6} e^{-0.001x - 0.002y} dy \right) dx \\ &= 6 \times 10^{-6} \int_0^{\infty} \left( \int_x^{\infty} e^{-0.002y} dy \right) e^{-0.001x} dx \\ &= 6 \times 10^{-6} \int_0^{\infty} \left( \frac{e^{-0.002x}}{0.002} \right) e^{-0.001x} dx \\ &= 0.003 \left( \int_0^{\infty} e^{-0.003x} dx \right) = 0.003 \left( \frac{1}{0.003} \right) = 1\end{aligned}$$

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**Figure 5-8** The joint probability density function of  $X$  and  $Y$  is nonzero over the shaded region.

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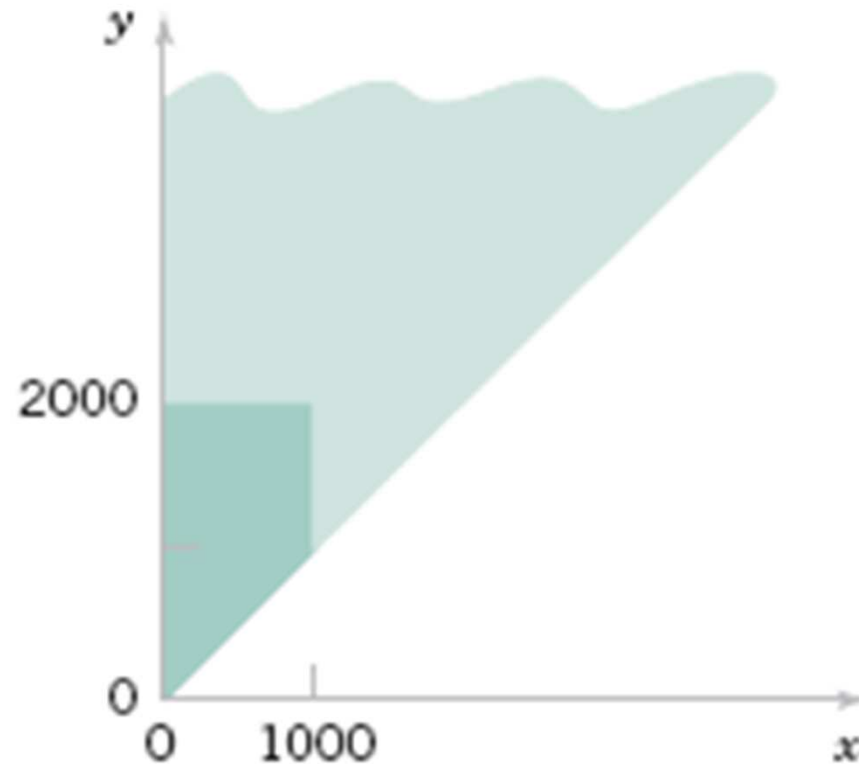
### Example 5-12

The probability that  $X < 1000$  and  $Y < 2000$  is determined as the integral over the darkly shaded region in Fig. 5-9.

$$\begin{aligned} P(X \leq 1000, Y \leq 2000) &= \int_0^{1000} \int_x^{2000} f_{XY}(x, y) dy dx \\ &= 6 \times 10^{-6} \int_0^{1000} \left( \int_x^{2000} e^{-0.002y} dy \right) e^{-0.001x} dx \\ &= 6 \times 10^{-6} \int_0^{1000} \left( \frac{e^{-0.002x} - e^{-4}}{0.002} \right) e^{-0.001x} dx \\ &= 0.003 \int_0^{1000} e^{-0.003x} - e^{-4} e^{-0.001x} dx \\ &= 0.003 \left[ \left( \frac{1 - e^{-3}}{0.003} \right) - e^{-4} \left( \frac{1 - e^{-1}}{0.001} \right) \right] \\ &= 0.003(316.738 - 11.578) = 0.915 \end{aligned}$$

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**Figure 5-9** Region of integration for the probability that  $X < 1000$  and  $Y < 2000$  is darkly shaded.

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## 5-2.2 Marginal Probability Distributions

### Definition

If the joint probability density function of continuous random variables  $X$  and  $Y$  is  $f_{XY}(x, y)$ , the **marginal probability density functions** of  $X$  and  $Y$  are

$$f_X(x) = \int_y f_{XY}(x, y) dy \quad \text{and} \quad f_Y(y) = \int_x f_{XY}(x, y) dx \quad (5-15)$$

where the first integral is over all points in the range of  $(X, Y)$  for which  $X = x$  and the second integral is over all points in the range of  $(X, Y)$  for which  $Y = y$



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### Example 5-13

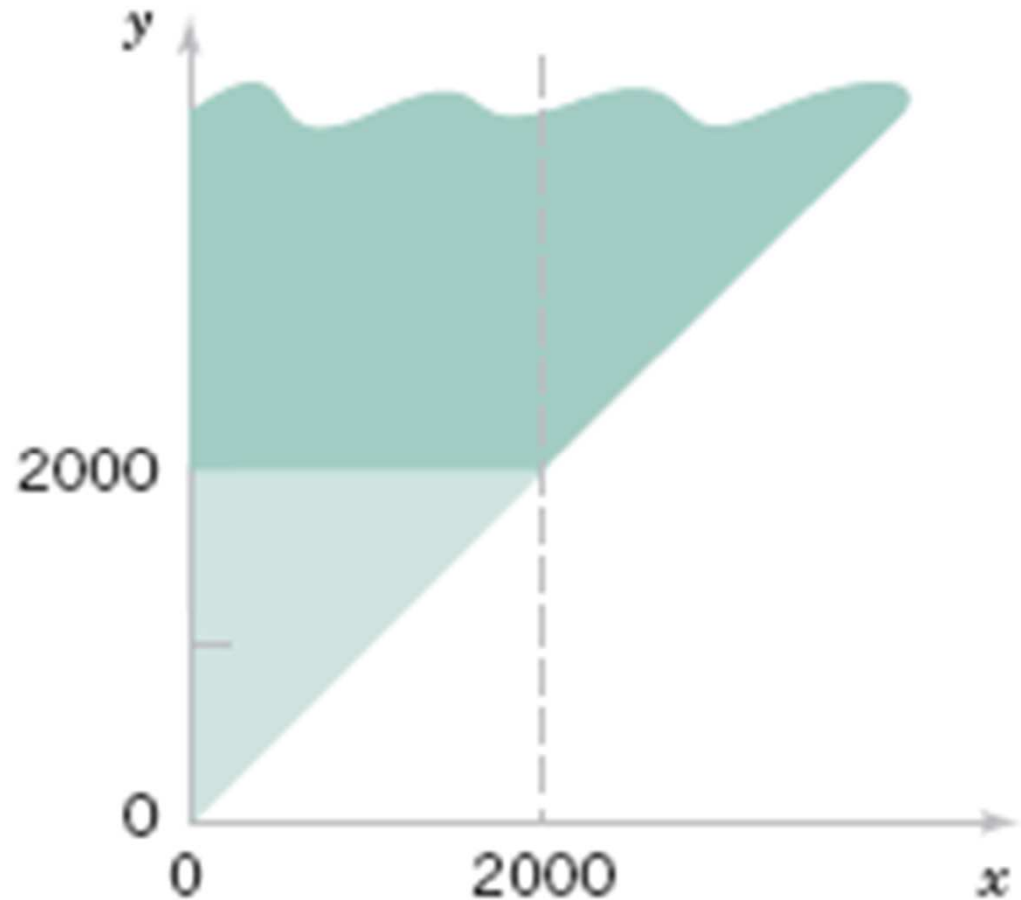
For the random variables that denote times in Example 5-12, calculate the probability that  $Y$  exceeds 2000 milliseconds.

This probability is determined as the integral of  $f_{XY}(x, y)$  over the darkly shaded region in Fig. 5-10. The region is partitioned into two parts and different limits of integration are determined for each part.

$$P(Y > 2000) = \int_0^{2000} \left( \int_{2000}^{\infty} 6 \times 10^{-6} e^{-0.001x - 0.002y} dy \right) dx \\ + \int_{2000}^{\infty} \left( \int_x^{\infty} 6 \times 10^{-6} e^{-0.001x - 0.002y} dy \right) dx$$

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**Figure 5-10** Region of integration for the probability that  $Y < 2000$  is darkly shaded and it is partitioned into two regions with  $x < 2000$  and  $x > 2000$ .



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### Example 5-13

The first integral is

$$\begin{aligned} 6 \times 10^{-6} \int_0^{2000} \left( \frac{e^{-0.002y}}{-0.002} \Big|_{2000}^{\infty} \right) e^{-0.001x} dx &= \frac{6 \times 10^{-6}}{0.002} e^{-4} \int_0^{2000} e^{-0.001x} dx \\ &= \frac{6 \times 10^{-6}}{0.002} e^{-4} \left( \frac{1 - e^{-2}}{0.001} \right) = 0.0475 \end{aligned}$$

The second integral is

$$\begin{aligned} 6 \times 10^{-6} \int_{2000}^{\infty} \left( \frac{e^{-0.002y}}{-0.002} \Big|_x^{\infty} \right) e^{-0.001x} dx &= \frac{6 \times 10^{-6}}{0.002} \int_{2000}^{\infty} e^{-0.003x} dx \\ &= \frac{6 \times 10^{-6}}{0.002} \left( \frac{e^{-6}}{0.003} \right) = 0.0025 \end{aligned}$$

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### Example 5-13

Therefore,

$$P(Y > 2000) = 0.0475 + 0.0025 = 0.05.$$

Alternatively, the probability can be calculated from the marginal probability distribution of  $Y$  as follows. For  $y > 0$

$$\begin{aligned} f_Y(y) &= \int_0^y 6 \times 10^{-6} e^{-0.001x - 0.002y} dx = 6 \times 10^{-6} e^{-0.002y} \int_0^y e^{-0.001x} dx \\ &= 6 \times 10^{-6} e^{-0.002y} \left( \frac{e^{-0.001x}}{-0.001} \Big|_0^y \right) = 6 \times 10^{-6} e^{-0.002y} \left( \frac{1 - e^{-0.001y}}{0.001} \right) \\ &= 6 \times 10^{-3} e^{-0.002y} (1 - e^{-0.001y}) \quad \text{for } y > 0 \end{aligned}$$

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### Example 5-13

We have obtained the marginal probability density function of  $Y$ . Now,

$$\begin{aligned} P(Y > 2000) &= 6 \times 10^{-3} \int_{2000}^{\infty} e^{-0.002y} (1 - e^{-0.001y}) dy \\ &= 6 \times 10^{-3} \left[ \left( \frac{e^{-0.002y}}{-0.002} \right) \Big|_{2000}^{\infty} - \left( \frac{e^{-0.003y}}{-0.003} \right) \Big|_{2000}^{\infty} \right] \\ &= 6 \times 10^{-3} \left[ \frac{e^{-4}}{0.002} - \frac{e^{-6}}{0.003} \right] = 0.05 \end{aligned}$$

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### 5-2.3 Conditional Probability Distributions

#### Definition

Given continuous random variables  $X$  and  $Y$  with joint probability density function  $f_{XY}(x, y)$ , the **conditional probability density function** of  $Y$  given  $X = x$  is

$$f_{Y|x}(y) = \frac{f_{XY}(x, y)}{f_X(x)} \quad \text{for} \quad f_X(x) > 0 \quad (5-16)$$

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### 5-2.3 Conditional Probability Distributions

Because the conditional probability density function  $f_{Y|x}(y)$  is a probability density function for all  $y$  in  $R_x$ , the following properties are satisfied:

$$(1) \quad f_{Y|x}(y) \geq 0$$

$$(2) \quad \int_{R_x} f_{Y|x}(y) dy = 1$$

$$(3) \quad P(Y \in B | X = x) = \int_B f_{Y|x}(y) dy \quad \text{for any set } B \text{ in the range of } Y$$

(5-17)

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### Example 5-14

For the random variables that denote times in Example 5-12, determine the conditional probability density function for  $Y$  given that  $X = x$ .

First the marginal density function of  $x$  is determined. For  $x > 0$

$$\begin{aligned} f_X(x) &= \int_x^{\infty} 6 \times 10^{-6} e^{-0.001x - 0.002y} dy = 6 \times 10^{-6} e^{-0.001x} \left( \frac{e^{-0.002y}}{-0.002} \bigg|_x^{\infty} \right) \\ &= 6 \times 10^{-6} e^{-0.001x} \left( \frac{e^{-0.002x}}{0.002} \right) = 0.003 e^{-0.003x} \quad \text{for } x > 0 \end{aligned}$$

This is an exponential distribution with  $\lambda = 0.003$ . Now, for  $0 < x$  and  $x < y$  the conditional probability density function is

$$\begin{aligned} f_{Y|x}(y) &= f_{XY}(x, y) / f_X(x) = \frac{6 \times 10^{-6} e^{-0.001x - 0.002y}}{0.003 e^{-0.003x}} \\ &= 0.002 e^{0.002x - 0.002y} \quad \text{for } 0 < x \quad \text{and} \quad x < y \end{aligned}$$

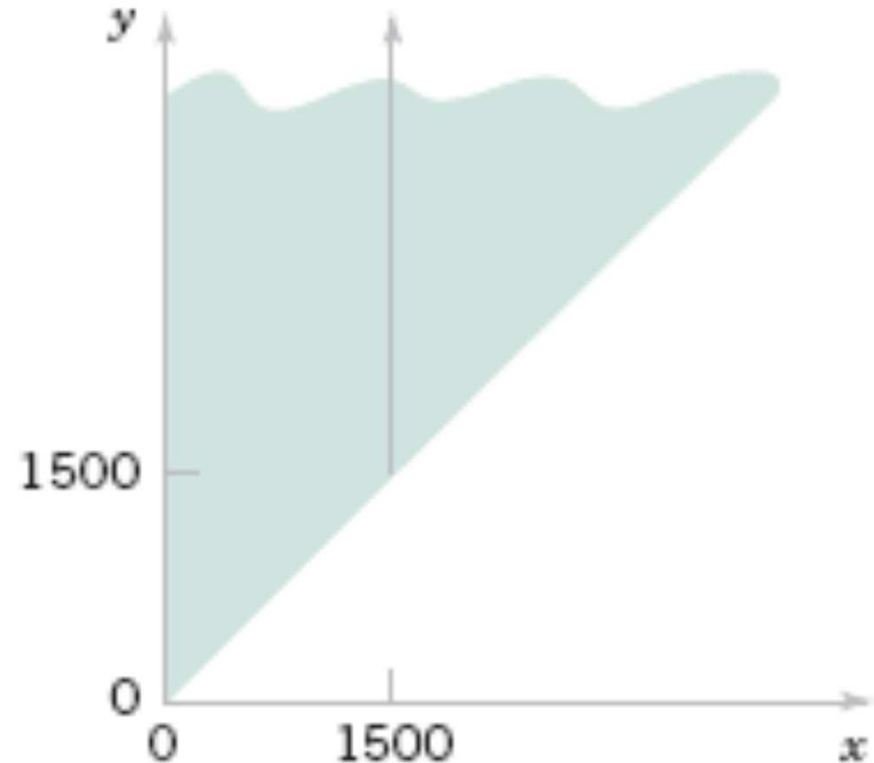


## 5-2 Two Continuous Random Variables

### Example 5-14

The conditional probability density function of  $Y$ , given that  $x = 1500$ , is nonzero on the solid line in Fig. 5-11.

**Figure 5-11** The conditional probability density function for  $Y$ , given that  $x = 1500$ , is nonzero over the solid line.



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### Definition: Conditional Mean and Variance

The **conditional mean** of  $Y$  given  $X = x$ , denoted as  $E(Y|x)$  or  $\mu_{Y|x}$ , is

$$E(Y|x) = \int y f_{Y|x}(y) dy$$

and the **conditional variance** of  $Y$  given  $X = x$ , denoted as  $V(Y|x)$  or  $\sigma_{Y|x}^2$ , is

$$V(Y|x) = \int_{-\infty}^{\infty} (y - \mu_{Y|x})^2 f_{Y|x}(y) dy = \int y^2 f_{Y|x}(y) dy - \mu_{Y|x}^2 \quad (5-18)$$

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## 5-2.4 Independence

### Definition

For continuous random variables  $X$  and  $Y$ , if any one of the following properties is true, the others are also true, and  $X$  and  $Y$  are said to be **independent**.

- (1)  $f_{XY}(x, y) = f_X(x)f_Y(y)$  for all  $x$  and  $y$
- (2)  $f_{Y|x}(y) = f_Y(y)$  for all  $x$  and  $y$  with  $f_X(x) > 0$
- (3)  $f_{X|y}(x) = f_X(x)$  for all  $x$  and  $y$  with  $f_Y(y) > 0$
- (4)  $P(X \in A, Y \in B) = P(X \in A)P(Y \in B)$  for any sets  $A$  and  $B$  in the range of  $X$  and  $Y$ , respectively. (5-19)

## **5-2 Two Continuous Random Variables**

### **Example 5-16**

For the joint distribution of times in Example 5-12, the

- Marginal distribution of  $Y$  was determined in Example 5-13.
- Conditional distribution of  $Y$  given  $X = x$  was determined in Example 5-14.

Because the marginal and conditional probability densities are not the same for all values of  $x$ , property (2) of Equation 5-18 implies that the random variables are not independent. The fact that these variables are not independent can be determined quickly by noticing that the range of  $(X, Y)$ , shown in Fig. 5-8, is not rectangular. Consequently, knowledge of  $X$  changes the interval of values for  $Y$  that receives nonzero probability.

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### Example 5-18

Let the random variables  $X$  and  $Y$  denote the lengths of two dimensions of a machined part, respectively. Assume that  $X$  and  $Y$  are independent random variables, and further assume that the distribution of  $X$  is normal with mean 10.5 millimeters and variance  $0.0025$  (millimeter)<sup>2</sup> and that the distribution of  $Y$  is normal with mean 3.2 millimeters and variance  $0.0036$  (millimeter)<sup>2</sup>. Determine the probability that  $10.4 < X < 10.6$  and  $3.15 < Y < 3.25$ .

Because  $X$  and  $Y$  are independent,

$$\begin{aligned} P(10.4 < X < 10.6, 3.15 < Y < 3.25) &= P(10.4 < X < 10.6)P(3.15 < Y < 3.25) \\ &= P\left(\frac{10.4 - 10.5}{0.05} < Z < \frac{10.6 - 10.5}{0.05}\right) P\left(\frac{3.15 - 3.2}{0.06} < Z < \frac{3.25 - 3.2}{0.06}\right) \\ &= P(-2 < Z < 2)P(-0.833 < Z < 0.833) = 0.566 \end{aligned}$$

where  $Z$  denotes a standard normal random variable.

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### Example 5-20

In an electronic assembly, let the random variables  $X_1, X_2, X_3, X_4$  denote the lifetimes of four components in hours. Suppose that the joint probability density function of these variables is

$$f_{X_1 X_2 X_3 X_4}(x_1, x_2, x_3, x_4) = 9 \times 10^{-2} e^{-0.001x_1 - 0.002x_2 - 0.0015x_3 - 0.003x_4}$$
$$\text{for } x_1 \geq 0, x_2 \geq 0, x_3 \geq 0, x_4 \geq 0$$

What is the probability that the device operates for more than 1000 hours without any failures?

The requested probability is  $P(X_1 > 1000, X_2 > 1000, X_3 > 1000, X_4 > 1000)$ , which equals the multiple integral of  $f_{X_1 X_2 X_3 X_4}(x_1, x_2, x_3, x_4)$  over the region  $x_1 > 1000, x_2 > 1000, x_3 > 1000, x_4 > 1000$ . The joint probability density function can be written as a product of exponential functions, and each integral is the simple integral of an exponential function. Therefore,

$$P(X_1 > 1000, X_2 > 1000, X_3 > 1000, X_4 > 1000) = e^{-1-2-1.5-3} = 0.00055$$

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### Definition: Marginal Probability Density Function

If the joint probability density function of continuous random variables  $X_1, X_2, \dots, X_p$  is  $f_{X_1 X_2 \dots X_p}(x_1, x_2, \dots, x_p)$ , the **marginal probability density function** of  $X_i$  is

$$f_{X_i}(x_i) = \int \int \dots \int_{R_{x_i}} f_{X_1 X_2 \dots X_p}(x_1, x_2, \dots, x_p) dx_1 dx_2 \dots dx_{i-1} dx_{i+1} \dots dx_p \quad (5-21)$$

where the integral is over all points in the range of  $X_1, X_2, \dots, X_p$  for which  $X_i = x_i$ .

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### Mean and Variance from Joint Distribution

$$E(X_i) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \dots \int_{-\infty}^{\infty} x_i f_{X_1 X_2 \dots X_p}(x_1, x_2, \dots, x_p) dx_1 dx_2 \dots dx_p$$

and

(5-22)

$$V(X_i) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \dots \int_{-\infty}^{\infty} (x_i - \mu_{X_i})^2 f_{X_1 X_2 \dots X_p}(x_1, x_2, \dots, x_p) dx_1 dx_2 \dots dx_p$$



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### Distribution of a Subset of Random Variables

If the joint probability density function of continuous random variables  $X_1, X_2, \dots, X_p$  is  $f_{X_1 X_2 \dots X_p}(x_1, x_2, \dots, x_p)$ , the **probability density function** of  $X_1, X_2, \dots, X_k, k < p$ , is

$$\begin{aligned} f_{X_1 X_2 \dots X_k}(x_1, x_2, \dots, x_k) \\ = \int_{R_{x_{k+1}}} \int_{R_{x_{k+2}}} \dots \int f_{X_1 X_2 \dots X_p}(x_1, x_2, \dots, x_p) dx_{k+1} dx_{k+2} \dots dx_p \end{aligned} \quad (5-23)$$

where the integral is over all points in the range of  $X_1, X_2, \dots, X_k$  for which  $X_1 = x_1, X_2 = x_2, \dots, X_k = x_k$ .

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### Conditional Probability Distribution

#### Definition

Continuous random variables  $X_1, X_2, \dots, X_p$  are independent if and only if

$$f_{X_1 X_2 \dots X_p}(x_1, x_2, \dots, x_p) = f_{X_1}(x_1) f_{X_2}(x_2) \dots f_{X_p}(x_p) \quad \text{for all } x_1, x_2, \dots, x_p \quad (5-24)$$

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### Example 5-23

Suppose  $X_1$ ,  $X_2$ , and  $X_3$  represent the thickness in micrometers of a substrate, an active layer, and a coating layer of a chemical product. Assume that  $X_1$ ,  $X_2$ , and  $X_3$  are independent and normally distributed with  $\mu_1 = 10000$ ,  $\mu_2 = 1000$ ,  $\mu_3 = 80$ ,  $\sigma_1 = 250$ ,  $\sigma_2 = 20$ , and  $\sigma_3 = 4$ , respectively.

The specifications for the thickness of the substrate, active layer, and coating layer are  $9200 < x_1 < 10800$ ,  $950 < x_2 < 1050$ , and  $75 < x_3 < 85$ , respectively. What proportion of chemical products meets all thickness specifications? Which one of the three thicknesses has the least probability of meeting specifications?

The requested probability is  $P(9200 < X_1 < 10800, 950 < X_2 < 1050, 75 < X_3 < 85)$ . Because the random variables are independent,

$$\begin{aligned} &P(9200 < X_1 < 10800, 950 < X_2 < 1050, 75 < X_3 < 85) \\ &= P(9200 < X_1 < 10800)P(950 < X_2 < 1050)P(75 < X_3 < 85) \end{aligned}$$

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### Example 5-23

After standardizing, the above equals

$$P(-3.2 < Z < 3.2)P(-2.5 < Z < 2.5)P(-1.25 < Z < 1.25)$$

where  $Z$  is a standard normal random variable. From the table of the standard normal distribution, the above equals

$$(0.99862)(0.98758)(0.78870) = 0.7778$$

The thickness of the coating layer has the least probability of meeting specifications. Consequently, a priority should be to reduce variability in this part of the process.